

# R300 – Advanced Econometric Methods

## PROBLEM SET 1 - SOLUTIONS

Due by Mon. October 12

1. Let  $y$  and  $x$  be scalar random variables. Show that the following hold.

(i)  $\text{cov}(E(y|x), y - E(y|x)) = 0$ .

(ii)  $\text{var}(y - E(y|x)) \leq \text{var}(y)$ .

(iii)  $\text{var}(y - E(y|x)) = \text{var}(y)$  when  $y$  and  $x$  are independent.

(i) Let  $\varepsilon = y - E(y|x)$ . Then we need to show that

$$\text{cov}(E(y|x), \varepsilon) = 0.$$

By definition of the covariance in the first step, the fact that  $E(\varepsilon) = 0$  in the second step, and the law of iterated expectations and the fact that  $E(\varepsilon|x) = 0$  in the third and fourth step, we have

$$\begin{aligned}\text{cov}(E(y|x), \varepsilon) &= E(E(y|x)\varepsilon) - E(E(y|x))E(\varepsilon) \\ &= E(E(y|x)\varepsilon) \\ &= E(E(y|x)E(\varepsilon|x)) \\ &= 0.\end{aligned}$$

(ii) We have, using the law of total variance and the fact that  $E(\varepsilon|x) = 0$  (and so has zero variance)

$$\begin{aligned}\text{var}(\varepsilon) &= \text{var}(E(\varepsilon|x)) + E(\text{var}(\varepsilon|x)) \\ &= E(\text{var}(\varepsilon|x)) \\ &= E(\text{var}(y|x)) \\ &\leq E(\text{var}(y|x)) + \text{var}(E(y|x)) \\ &= \text{var}(y)\end{aligned}$$

(iii) This is immediate from the solution to (ii). By independence,  $E(y|x) = E(y)$  and so  $\text{var}(E(y|x)) = 0$  from which the equality follows.

---

2. Show that the following functions are probability mass/density functions and compute their first two moments.

(i)  $f(x) = ax^{a-1}$ ,  $x \in (0, 1)$  and  $a > 0$ .

(ii)  $f(x) = n^{-1}$ ,  $x = 1, 2, \dots, n$  for  $n$  a positive integer.

(iii)  $f(x) = (3/2)(x - 1)^2$ ,  $x \in (0, 2)$ .

---

(i) It suffices to show that  $f$  is non-negative and integrates to one. Here,  $f$  is clearly non-negative as  $a > 0$  and  $x$  is positive. Further,

$$\int_0^1 f(x) dx = a \int_0^1 x^{a-1} dx = x^a \Big|_0^1 = 1.$$

So  $f$  is a density function.

Next, the first two moments are

$$\int_0^1 x (a x^{a-1}) dx = a \int_0^1 x^a dx = \frac{a}{1+a}, \quad \int_0^1 x^2 (a x^{a-1}) dx = \frac{a}{2+a},$$

by straightforward integration.

(ii) Here,  $f$  is the uniform mass function on the interval  $[1, 2, \dots, n]$ , i.e., it places mass  $n^{-1}$  on each of the  $n$  masspoints. So, clearly, mass is positive and sums up to one over the support.

The first two moments are

$$\sum_{i=1}^n i/n = (n+1)/2, \quad \sum_{i=1}^n i^2/n = (n+1)(2n+1)/6.$$

(iii) That  $f$  is a pdf is easy to see. The integral is straightforward as  $f$  is polynomial (as in (i)),

$$\frac{3}{2} \int_0^2 (x^2 + 1 - 2x) dx = \frac{3}{2} \left( \frac{x^3}{3} + x - x^2 \right) \Big|_0^2 = 1.$$

The first and second moments are 1 and  $8/5$ , respectively.

---

3. Suppose that, for a scalar random variable  $x$ ,

$$F(v) = P(x \leq v)$$

is continuous and strictly increasing in  $v$ . Define the random variable  $y = F(x)$ . Derive the distribution of  $y$ .

---

Let  $G(v) = P(y \leq v)$ . Then, for any  $v \in [0, 1]$ , we have

$$G(v) = P(F(x) \leq v) = P(x \leq F^{-1}(v)) = F(F^{-1}(v)) = v.$$

So,

$$G(v) = \begin{cases} 0 & \text{if } v < 0 \\ v & \text{if } v \in [0, 1] \\ 1 & \text{if } v \geq 1 \end{cases} .$$

This is the uniform distribution on the interval  $[0, 1]$ .

---

4. For each of the cases below, show that the score is mean zero and derive the efficiency bound for  $\theta$ .

- (i) Binomial with parameters  $(n, \theta)$ .
  - (ii) Exponential with mean  $1/\theta$ .
  - (iii) Exponential with mean  $\theta$ .
- 

(i) The sum of  $m$  independent Bernoulli random variables, each with success probability  $\theta$ , is Binomial  $(m, \theta)$ . The mass function at  $x = 0, 1, 2 \dots$  equals

$$f_{\theta}(x) = \binom{m}{x} \theta^x (1 - \theta)^{m-x}.$$

So,

$$\log f_{\theta}(x) = \log \binom{m}{x} + x \log \theta + (m - x) \log(1 - \theta).$$

Therefore,

$$\frac{\partial \log f_{\theta}(x)}{\partial \theta} = \frac{x}{\theta} - \frac{m - x}{1 - \theta} = \frac{x - m\theta}{\theta(1 - \theta)}$$

and

$$\frac{\partial^2 \log f_{\theta}(x)}{\partial \theta^2} = -\frac{x}{\theta^2} - \frac{m - x}{(1 - \theta)^2}.$$

Now,  $E_\theta(x) = m\theta$  and  $\text{var}_\theta(x) = m\theta(1-\theta)$ . Note that the score has mean zero. Its variance is

$$E_\theta \left( \frac{\partial \log f_\theta(x)}{\partial \theta} \frac{\partial \log f_\theta(x)}{\partial \theta} \right) = \frac{\text{var}_\theta(x)}{\theta^2(1-\theta)^2} = \frac{m}{\theta(1-\theta)}.$$

Equivalently,

$$-E_\theta \left( \frac{\partial^2 \log f_\theta(x)}{\partial \theta^2} \right) = E_\theta \left( \frac{x}{\theta^2} + \frac{m-x}{(1-\theta)^2} \right) = \frac{m}{\theta} + \frac{m}{1-\theta} = \frac{m}{\theta(1-\theta)}.$$

The efficiency bound then is  $\theta(1-\theta)/(nm)$ .

(ii) The exponential has

$$f_\theta(x) = \theta e^{-x\theta}$$

for  $x \geq 0$ . Here,

$$\log f_\theta(x) = \log \theta - x\theta$$

and so

$$\frac{\partial \log f_\theta(x)}{\partial \theta} = \frac{1}{\theta} - x, \quad \frac{\partial^2 \log f_\theta(x)}{\partial \theta^2} = -\frac{1}{\theta^2}.$$

The efficiency bound is  $\theta^2/n$ .

(iii) In this parametrization the exponential has

$$f_\theta(x) = \frac{1}{\theta} e^{-x/\theta}$$

for  $x \geq 0$ . Here,

$$\log f_\theta(x) = -\log \theta - \frac{x}{\theta}.$$

Thus,

$$\frac{\partial \log f_\theta(x)}{\partial \theta} = -\frac{1}{\theta} + \frac{x}{\theta^2} \quad \frac{\partial^2 \log f_\theta(x)}{\partial \theta^2} = \frac{1}{\theta^2} - \frac{2x}{\theta^3}.$$

The information becomes

$$1/\theta^2$$

The efficiency bound is  $\theta^2/n$ .

---